# Socio-Economic Risk Profiling of Child Nutritional Status: Evidence from the Most Populous Province of Pakistan

Abdu R Rahman<sup>1\*</sup>, Zahid Asghar<sup>2</sup>, Salman A Cheem<sup>3</sup>, Tahir Munir<sup>4</sup>

<sup>1</sup>Department of Global Health and Development, The Aga Khan University, Karachi, Pakistan <sup>2</sup>Department of Economics, Quaid-i-Azam University Islamabad, Islamabad, Pakistan <sup>3</sup>Department of Applied Sciences, National Textile University, Faisalabad, Pakistan <sup>4</sup>Department of Anaesthesiology, The Aga Khan University, Karachi, Pakistan

#### **Research Article**

### ABSTRACT

Received: 07-Jan-2023, Manuscript No. JSMS-23-86138; Editor assigned: 09-Jan-2023, Pre QC No. JSMS-23-86138 (PQ); Reviewed: 23-Jan-2023, QC No. JSMS-23-86138; Revised: 18-Apr-2023, Manuscript No. JSMS-23-86138 (R); Published: 01-Jun-2023, DOI: 10.4172/JSMS.9.2.011

**\*For Correspondence :** Abdu R Rahman, Department of Global Health and Development, The Aga Khan University, Karachi, Pakistan;

Email: abdur.rahman@stat.qau.edu.pk Citation: Rahman AR, et al. Socio-Economic Risk Profiling of Child Nutritional Status: Evidence from the Most Populous Province of Pakistan. RRJ Stats Math Sci. 2023;09:011.

**Copyright:** © 2023 Rahman AR, et al. This is an open-access article distributed under the terms of the Creative Commons Attribution License, which permits unrestricted use, distribution and reproduction in any medium, provided the original author and source are credited. Child malnutrition is considered one of the most focused areas of development, globally. The situation in developing countries, with lower literacy rate and lesser health awareness, is even more alarming and unpleasant. Pakistan a country of 220 million populations with literacy rate of 65 percent remains a prime candidate, worth studying the diverse nature of the issue. This research focuses on the analysis of Punjab based data the most populated province of Pakistan, sharing 50 percent of the total population of the country. Principally, this research advances the existing literature mainly on two fronts. Firstly, we study children nourishment status through ordinal scale and thus identify the more vulnerable and priority groups existent in the population. Secondly, we propose the use of WHO Infant and Young Children Feeding guidelines (IYCF) for food quality, as a determinant of child nourishment status. Also, we consider weight for age, as a composite anthropometric indicator to quantify the nourishment status of children of age under five years. Based on this indicator, child nourishment status is then categorized into three distinctive and hierarchical groups: severely malnourished (<-3.0 Z-score), moderately malnourished (-3.0 to -2.01 Z-score) and not malnourished ( $\geq 2.0$  Z score). The objectives are achieved by using the Multiple Indicator Cluster Survey (MICS) 2017-2018 data for the Punjab province comprehending a sample of 25211 children. We observe that 7% children can be ranked as severely malnourished whereas, 14.5% children stayed in the moderately malnourished category. Moreover, bivariate analysis reveals statistically significant association between children nourishment status and, food intake diversity, mother education and health awareness, child previous health history and economic status of the household. The explanatory power of the determinants of malnourishment is assessed by employing various modeling strategies capable of entertaining diverse ordinal structures. We use Proportional Odds Model (POM), Non-Proportional Odds Model (NPOM). Based on keen application of statistical modeling techniques, our study suggests that NPOM can be considered as a more sophisticated approach to explore the factors affecting the child malnutrition. The findings of this research imply that, government and development organizations need to focus, not only, on improvement of overall household well-being but also required advocating the urgency for balanced food.

**Keywords:** Child nourishment status; Food diversity; Generalized linear model; Multiple indicator cluster survey; Ordinal data

### INTRODUCTION

Nourishment status of individuals remains one of the elementary indicators delineating the extent of overall social and physical human welfare. The study of child nourishment status and its inter linkage with socio-economic growth has remained a core topic of various studies from diverse research areas. For example, Carlson suggested the overall nourishment status as a determinant of social capital. Hoddinott, et al., highlighted economic aspects of better nourishment by identifying lower health care expenditures as a natural outcome <sup>[1]</sup>. Similarly, Victora, et al., studied long term effects of malnutrition in children and listed, impaired cognitive development, lower attained schooling and increased risk of chronic diseases, as irreversible consequences of poor nutrition. Moreover, the targets 2.1 and 2.2 of Sustainable Development Goals (SDGs) focus to end hunger and reduce malnutrition to get rapid economic growth. For more recent accounts in this regard, see also Gul and Kibria; Talukder; Paul, et al. Despite being globally focused area of improvement, according to UNICEF report about 2.6 million people die each year due to malnutrition, see also Black, et al. Additionally, the number of malnourished people increased from 785.8 million in 2014 to 821.6 million in 2018, worldwide. The situation among more vulnerable and fragile group, that is, the children under five years age, is even more upsetting. According to the world health organization report, 155 million children under five years of age were stunted (too short with respect to age). Whereas 52 million children were considered as wasted (too thin with respect to height) and 462 million were assessed as underweight (small weight with respect to age). The South Asian region is doomed to be the major contributor in the odds of malnourished children in the world. This region alone contributes to 40% (59.4 million) of world's stunted children and 53% (27 million) of world's wasted children (UNICEF, 2018). Within this region, the standing of Pakistan is even more unpleasant Pakistan was observed as highest shareholder in stunting with 45% in 2012-2013. Although, the percentage of malnourished children has declined over the past 15 years, but still 40.2%, 17.7% and 28.9% of young children are stunted, wasted and underweight, respectively, see also, Mushtaq, et al; Batool, et al; Gul and Kibria.

Motivated by the relevant nature of the issue, this study focuses on exploring the potential factors determining the child nourishment status in Punjab province of Pakistan. The objectives are achieved by a detailed analysis of Multiple Indicator Cluster Survey (MICS) 2017-2018 data for Punjab. The significance of the analysis of Punjab based data remains intact as the province shares half of the total population of the country. Thus, a small percentage of malnourished children add a large absolute number to the total of malnourished population. The goals of this research are two fold. Firstly, to devise more sophisticated strategies capable of entertaining the underlying ordinal structure, prevalent in the nourishment status. Secondly, based on the findings, facilitate the meaningful and better informed policies to achieve targets 2.1 and 2.2 of the SDGs. In per suite of the afore mentioned goals, along with detailed exploratory analysis, we employ variety of Generalized Linear Models (GLMs), while stratifying nourishment status into three hierarchical categories. Furthermore, along with routinely used literature driven factors, we propose the use of food diversity index, to enhance the understanding of the issue <sup>[2]</sup>.

This article is divided into four major sections. In next section, we provide a comprehensive account of data, covariates, their measurements, and modelling schemes employed in this research. Section 3 offers a detail discussion about the findings of the study with respect to various modeling themes subject to specific assumptions. Lastly, section 4 provides brief overview of the matter focusing on policy implications <sup>[3]</sup>.

#### MATERIALS AND METHODS

#### Data and design

The data for current study are extracted from 6<sup>th</sup> round of MICS Punjab a data collection exercise jointly conducted by Punjab Bureau of Statistics and UNICEF. The aim of MICS is to assemble information on socioeconomic indicators at individual level as well as household level and health awareness among women. Thus, MICS serves as primary tool in district level policy formulation in Punjab. The authenticity of the information collected through MICS is well celebrated among organizational and academic research circles. For MICS (2017-2018), two stage stratified sampling design was adopted. At first stage, each district of the province was stratified into rural and urban enumeration blocks. The weights to the enumeration blocks were assigned through probability proportional to size sampling scheme. In second stage, from each selected enumeration block, 20 households were finally approached as member of selected sample through systematic sampling scheme. Thus, as a result of MICS (2017-2018) 53,840 households were approached from 2692 enumeration blocks. Among the selected households, 51660 households were successfully interviewed with a response rate of 97.9 percent. At individual level, MICS (2017-2018) comprehends the information about 74010 women, 27094 men and 42408 children under the age of five years <sup>[4]</sup>.

#### Response variable

The child's nourishment status is usually estimated by using three contemporary anthropometric indicators, such as height for age, weight for height, and weight for age. Among three, the weight for age is commonly considered as the

most attractive representative of nourishment status, see for example, national institute of population research. The most convincing feature of the indicator remains it ability to cater weight for height and height for age indices, jointly, and therefore, successfully measures both chronic and acute malnutrition <sup>[5]</sup>. In this research we use weight-for-age indicator to assess child's overall nourishment status. Following the WHO guidelines, the indicator is measured as Z-score ranging from -6 to +6. After calculating Z-scores, we stratified nourishment status into three ordered categories, such as, (1)-not malnourished ( $\geq$  -2 Z-score), (2)-moderately malnourished (-3.0 to -2.01 Z-score) and (3)-Severely malnourished (<-3 Z-score). The analysis of the nourishment status is persuaded with respect to hierarchical categories <sup>[6]</sup>.

### Explanatory variables

Based on existing literature and keeping the objectives of the study in view, various demographic and socioeconomic factors are considered as determinants of the nourishment status of children. Table 1 below contains the details of the covariates included in this study along with the details of their coding schemes indicating the measuring process [7].

Sr. no	Covariates	Measurements		
1	Divisions	Bahawalpur=1, DG Khan=2, Faisalabad=3, Gujranwala=4, Lahore=5, Multan=6, Rawalpindi=7, Sahiwal=8, Sargodha=9		
Chil	d history			
2	Child age in months	0-5=1, 6-23=2, above 23=3		
3	Child ill with fever during the last two weeks before the interview	Yes=1, no=0		
4	Child had diarrhea during the last two weeks before the interview	Yes=1, no=0		
5	Food diversity	None=0, anyone=1, any two=2, any three=3, at least four=4		
6	Size of child at birth	Above average=1, average=2, below average=3		
Mot	her education and health awareness			
7	Mother education	None/pre-school=1, primary=2, middle=3, secondary=4, higher=5		
8	Child having mothers those who received antenatal care during pregnancy	Yes=1, no=0		
Household socio-economic status				
9	Household wealth quintile index	Poorest=1, poor=2, middle=3, rich=4, richest=5		
10	Household own agricultural land	Yes=1, no=0		

**Quantification of food diversity:** Over the last two decades variety of measures of food consumption diversity, focusing different dimensions of nutrition, have been defined. More recently, WHO proposed IYCF food diversity index based on genetic family membership of variety of food subgroups ensuring the food richness, see also Swindale and Bilinsky. Using the available information about the food consumption history of child in MICS (2017-2018) data, we arranged food items into seven major groups based on IYCF guidelines. These food groups are (1) roots, tubers and grains; (2) nuts and legumes; (3) flesh foods (poultry, liver/organ meats and fish); (4) eggs; (5) vegetables, vitamin A and rich fruits; (6) dairy products (yogurt, milk and cheese); (7) other vegetables and fruits. The value "0" is assigned to indicate the non-utilization of food item according to IYCF guidelines. Further, the food quality is ranked less diverse if food intake is supported through less than four groups and is assigned a value of "1". Whereas the food quality is considered as diversified and enriched, if a child is consuming food items from at least four groups <sup>[8]</sup>.

### Ordinal regression models

In this study, since response variable is measured on ordinal scale to identify the priority groups existent in the population therefore, we employ variety of modeling schemes capable of entertaining the underlying hierarchy. We studied Proportional Odds Model (POM), Non-Proportional Odds Model (NPOM), Unconstrained Partial Proportional Odd Models (UPPOM) and Constrained Partial Proportional Odd Models (CPPOM), in the search of most appropriate schemes to model the nutrition status of children <sup>[9]</sup>.

### Proportional Odds Model (POM)

The POM is conventionally used when the parallel line assumption is satisfied in ordered outcome variable, that is, if

rate of change remains constant among all categories of response variable, Bender and Grouven. In general, POM can be written as:

$$Prob(Y \le y_{j|z}) = \left[\frac{\exp(\lambda_j - z'\beta)}{1 + \exp(\lambda_j - z'\beta)}\right] \quad \text{ for } j = 1, 2, \dots, J - 1.$$

Where  $\underline{Y}$  is vector compiling the categories of outcome variable and z' represents the covariates associated with j'th category of outcome variable. Further,  $\beta = (\beta_1, \beta_2, ..., \beta_k)'$  denotes parameters associated with covariates. Moreover,  $\lambda_j$  represents the unknown parameter, such as  $\lambda_1 \leq \lambda_2 \leq \cdots \leq \lambda_{(j-1)}$  highlighting the ordered threshold <sup>[10-15]</sup>. On further simplification, we write the above given cumulative formation as follows,

$$\operatorname{logit} \left[ \frac{\operatorname{Prob}(Y \leq y_{j|z})}{\operatorname{Prob}(Y \geq y_{j|z})} \right] = \lambda_{j} - z'\beta.$$

#### Non-Proportional Odds Model (NPOM)

The NPOM models allows the relaxation of parallel lines assumption Fu, et al. The generalized model is expressed in the following equation:

$$Prob(Y \leq y_{j|z}) = \begin{bmatrix} \frac{exp(\tau_j - z'\beta_j)}{1 + exp(\tau_j - z'\beta_j)} \end{bmatrix} \quad \text{for} \quad j = 1, 2, \dots, J-1.$$

Where,  $\tau_j$  maintains the ordinal structure by allowing the estimation of non-equivalent odds across the cumulative categories of response variable. A more workable expression of the model is derived after some simplification, as follows,

$$\text{logit} \left[ \frac{\text{Prob}\left( \textbf{Y} \leq \textbf{y}_{j|z} \right)}{\text{Prob}\left( \textbf{Y} \geq \textbf{y}_{j|z} \right)} \right] = \tau_j - z' \beta.$$

#### Un-constrained Partial Proportional Odds Model (UPPOM)

The UPPOM are applicable when parallel lines assumption holds for some covariates but need to relax for others and thus provides more general working environment as compared to the aforementioned models, Ananth and Kleinbaum. The objective is achieved by defining two distinctive groups of coefficients, such as.

$$\text{Prob}\big(Y \leq y_{j|z}\big) = \begin{bmatrix} \frac{\exp(-\alpha_m - z'\beta - g'\gamma_m)}{1 + \exp(-\alpha_m - z'\beta - g'\gamma_m)} \end{bmatrix} \quad \text{ for } m = 1, 2, \dots, k.$$

Here,  $\underline{z}$  is a vector of covariates satisfying the parallel lines assumption, whereas  $\underline{g}$  is a vector containing the covariates not satisfying the parallel lines assumption. In the above model,  $\gamma_m$  explains the increment changes in logit for non-proportional variables. When the parallel line assumption is justified and the  $\gamma_m=0$ , the model is treated as POM, Peterson and Harrell Jr.

#### Constrained Partial Proportional Odds Model (CPPOM)

The CPPOM is transformed unconstrained model and capably reduces the number parameters involved in the estimation of overall model. The transformation is attained by multiplying the coefficients at varying cut-points by a predetermined fixed scalar, Ananth and Kleinbaum. A general mathematical form of the model is described below:

$$\text{Prob}\big(Y \leq y_{j|z}\big) = \begin{bmatrix} \frac{\exp\left(-\alpha_m - z'\beta - g'\gamma\Gamma_m\right)}{1 + \exp\left(-\alpha_m - z'\beta - g'\gamma\Gamma_m\right)} \end{bmatrix} \quad \text{ for } m = 1, 2, \dots, k.$$

Here,  $\Gamma_m$  is the predetermined scalar and g vector is q scaled, where cumulative logits are computed by multiplying  $\gamma$  and  $\Gamma_m$ .

### **RESULTS AND DISCUSSION**

#### **Exploratory analysis**

**Child history:** Table 2 comprehends the results of exploratory analysis of the MICS (2017-2018) data for Punjab. We observe a significant association between child age and their nutrition status. Children of smaller age groups are more likely suffer malnutrition as compared to higher age categories. These findings indicate that children of age 0-5 months are more fragile than other groups and thus become the priority group with respect to age. Further, there is statistically significant association between child suffering from fever in last two weeks and malnutrition. Also, diarrhea is witnessed to be significantly associated with nutrition status of children. A child suffered from diarrhea is more likely to face malnutrition than those who did not suffered <sup>[16]</sup>. Similarly, we found statistically significant association between child status of child. The children with below average size at birth

remained more vulnerable towards malnutrition as compared to children with average and above average size at birth. Overall, we may say that child health history emerged as significant determinant of child nutrition status.

Coverietes	Not malnourished		Moderate malnourished		Severely malnourished		Total		
Covariates	percent	frequency	percent	frequency	percent	frequency			
Child age in months									
0-5	73.8	3181	14.9	643	11.3	486	4310		
6-23	80.1	8574	13.4	1432	6.6	703	10709		
Above 23	78.9	8038	15.5	1577	5.7	577	10192		
Chi-squared P-v	alue (<0.0	01)							
Child ill with fever in last two weeks before interview									
No	79.2	14376	14.2	2574	6.6	1202	18152		
Yes	76.7	5417	15.3	1078	8	564	7059		
Chi-squared P-v	alue (<0.0	01)							
	Chil	d had diarrh	ea in last tv	vo weeks befor	e interview				
No	79.2	16829	14.1	2991	6.7	1424	21244		
Yes	74.7	2964	16.7	661	8.6	342	3967		
Chi-squared P-v	alue (<0.0	01)							
Chil	d having m	nothers those	e who receiv	ed antenatal (	are during	pregnancy			
No	69.4	2025	19.3	562	11.3	329	2916		
Yes	79.7	17768	13.9	3090	6.4	1437	22295		
Chi-squared P-value (<0.001)									
		Agı	riculture lan	d ownership					
No	77.2	13075	15.4	2615	7.3	1,243.00	16933		
Yes	81.2	6718	12.5	1037	6.3	523	8278		
Chi-squared P-v	alue (<0.0	01)							
			Mother ed	lucation					
None	71.8	7836	18.6	2026	9.6	1050	10912		
Primary	79	4130	14.7	770	6.2	326	5226		
Middle	83	2198	11.3	299	5.7	151	2648		
Secondary	86.5	2880	9.4	314	4.1	137	3331		
Higher	88.8	2749	7.9	243	3.3	102	3094		
Chi-squared P-v	alue (<0.0	01)							
			Food divers	sity score					
None	73.4	4196	15.1	861	11.6	663	5720		
Any one	77.9	2803	15.1	544	7	250	3597		
Any two	81	2261	13.1	367	5.9	164	2792		
Any three	84.3	1486	11.2	197	4.5	79	1762		
Atleast four	79.8	9047	14.8	1683	5.4	610	11340		
Chi-squared P-v	alue (<0.0	01)							
			Size of chil	d at birth	1	1			
Above average	85.1	1560	11	202	3.9	72	1834		
Average	79.9	15269	13.9	2648	6.2	1,189	19106		

Table 2. Bivariate analysis of children's nourishment status with respect to the explanatory factors.

Below average	69.4	2964	18.8	802	11.8	505	4271			
Chi-squared P-value (<0.001)										
Household wealth quintile index										
Poorest	68	4215	20.9	1296	11.1	688	6199			
Poor	75.4	4140	16.6	910	8	439	5489			
Middle	82.7	4428	11.9	636	5.5	292	5356			
Rich	84.4	3898	11.1	512	4.5	209	4619			
Richest	87.7	3112	8.4	298	3.9	138	3548			
Chi-squared P-value (<0.001)										
			Divis	ion						
Bahawalpur	70	1514	19.9	430	10.1	218	2162			
DG Khan	73	2152	18.4	542	8.6	254	2948			
Faisalabad	78.1	2445	15.1	474	6.8	212	3131			
Gujranwala	83.3	3438	11.5	475	5.1	212	4125			
Lahore	80	2530	13.8	436	6.2	196	3162			
Multan	75.8	2303	15	456	9.2	278	3037			
Rawalpindi	86.9	2172	8.2	204	4.9	123	2499			
Sahiwal	80.3	1358	14.3	241	5.4	92	1691			
Sargodha	76.6	1881	16	394	7.4	181	2456			
Chi-squared P-v	alue (<0.0	001)								
Total	78.5	19793	14.5	3652	7	1766	25211			

**Mother education and health awareness:** A significant association is witnessed between mother education and nutrition status of children. We observe a consistent decrease in the proportion of malnutrition children having mothers with higher education. Similar patterns in nutrition status are found with respect to mother health awareness. Children with mothers having received antenatal care during pregnancy are less likely to suffer malnutrition as compared to their other peers.

**Household socio-income status:** Household socio-economic standing is also found a significant associate of child nutrition status. Children of families with higher wealth quartile are less likely to expose to malnutrition. Also, ownership of agricultural land significantly associates with nutrition status of child.

### Statistical modelling

In this section, we further proceed our investigation through the employment of more appropriate regression strategies. First, we used POM to explore the children nourishment status with respect to aforementioned covariates. The Table 3 presents the results of POM.

Covariates	Estimates	Standard errors	Odds ratios	95% C.I (OR)			
Intercept1	0.151***	0.133					
Intercept2	1.494***	0.134					
Child's monthly age (Ab	ove 23 months as	reference)					
0-5	-0.237*	0.107	0.79*	0.641-0.972			
06-11	-0.351***	0.092	0.705***	0.589-0.843			
Child ill with fever in la	st two weeks (No	as reference)					
Yes	0.078*	0.036	1.081*	1.008-1.16			
Child had diarrhea in last two weeks (No as reference)							
Yes	0.124**	0.044	1.132**	1.04-1.232			
Received Antenatal Care (No as reference)							

Yes	-0.239***	0.047	0.788***	0.72-0.863					
Household has own agriculture land (No as reference)									
Yes	-0.237***	0.035	0.79***	0.738-0.846					
Mother's education (None as reference)									
Primary	-0.158***	0.043	0.854***	0.786-0.929					
Middle	-0.287***	0.062	0.752***	0.667-0.847					
Secondary	-0.455***	0.064	0.635***	0.561-0.718					
Higher	-0.577***	0.074	0.562***	0.487-0.649					
Food diversity score (N	one as reference)								
Any one	-0.301***	0.059	0.741***	0.66-0.832					
Any two	-0.377***	0.068	0.687***	0.602-0.784					
Any three	-0.512***	0.082	0.6***	0.512-0.704					
At least four	-0.667***	0.1	0.514***	0.423-0.625					
Size of child at birth (A	Size of child at birth (Above average as reference)								
Average	0.318***	0.07	1.375***	1.2-1.574					
Below average	0.816***	0.075	2.26***	1.952-2.616					
Wealth quintile (Poores	st as reference)								
Poor	-0.231***	0.044	0.795***	0.729-0.866					
Middle	-0.51***	0.053	0.601***	0.543-0.666					
Rich	-0.535***	0.061	0.587***	0.521-0.66					
Richest	-0.66***	0.076	0.518***	0.447-0.6					
Divisions (Bahawalpur	as reference)								
DG Khan	-0.231***	0.065	0.794***	0.7-0.902					
Faisalabad	-0.171*	0.066	0.844*	0.742-0.96					
Gujranwala	-0.269***	0.067	0.765***	0.671-0.871					
Lahore	-0.131	0.069	0.878	0.768-1.004					
Multan	-0.105	0.065	0.901	0.794-1.024					
Rawalpindi	-0.532***	0.08	0.588***	0.503-0.687					
Sahiwal	-0.396***	0.079	0.674***	0.577-0.786					
Sargodha	-0.22**	0.069	0.804**	0.703-0.919					
Goodness of fit (Likelik	nood ratio); $\chi^2$ (28	)=1456.22; P-Value=<0.0	001; Pseudo R-squa	red=0.044; n=25211					
*, ** and *** indicates	o<0.05, p<0.01 ar	nd p<0.001, respectively.							

As a pre-requisite to POM, before interpreting the results the legitimacy of parallel line assumption was tested through Brant test. We found that the basic assumption of the POM does not hold for the MICS (2017-2018) Punjab based data. The Table 4 presents the results of Brant test.

Table 4. Dialit lest results.						
Osussistas	Chi-	Dualuas				
Covariates	Squarea	P-values				
Child's age	3.78	0.052				
Child ill with fever	1.22	0.27				
Child had diarrhea	0.69	0.405				
Received antenatal care	3.93	0.047				
Own agriculture land	3.4	0.065				
Mother education	0.09	0.758				
Food diversity score	9.7	0.002				
Size of child at birth	10.41	0.001				
Wealth quintile	0.24	0.628				
Divisions	2.21	0.138				
Overall	106.46	<0.0001				

Table 4. Brant test results

Due to the violation of parallel line assumption, we proceed by using NPOM as an alternative to the POM. The Table 5 compiles the outcomes. NPOM results shows that the odds ratios associated with the variables child's age, mother education and wealth index quintile vary for each category of the response.

Iable 5. The results of NPOM.									
Comparisons									
	Not malnourished vs. (Severely and moderately (Not and moderately malnourished) vs. sever								
		malnou	rished)			malnou	rished		
Covariates	Estimat	Standard	Odds	95% C.I	Estimates	Standard	Odds	95% C.I	
	es	errors	ratios	(OR)		errors	ratios	(OR)	
Intercept	-0.17	0.134			-1.38***	0.228			
Child's mont	hly age (ab	ove 23 as refere	ence)						
0 - 5	-0.254*	0.108	0.777*	0.63- 0.958	-0.208	0.185	0.813	0.567- 1.166	
06-11	-0.359***	0.093	0.7	0.584- 0.838	-0.365*	0.168	0.695*	0.501- 0.964	
Child ill with	fever in la	st two weeks (N	o as referen	ce)					
Yes	0.073*	0.037	1.075*	1.001- 1.154	0.137*	0.057	1.147*	1.026- 1.282	
Child had dia	arrhea in Ia	st two weeks (N	o as referen	ce)					
Yes	0.129**	0.044	1.137**	1.044- 1.239	0.092	0.068	1.097	0.961- 1.251	
Received an	tenatal car	e (No as referen	ce)						
Yes	-0.224***	0.047	0.8***	0.73- 0.877	-0.344***	0.069	0.71***	0.621- 0.813	
Household h	as own agr	iculture land (N	o as referen	ce)					
Yes	-0.246***	0.036	0.783***	0.731- 0.839	-0.164**	0.056	0.849**	0.762- 0.946	
Mother's edu	ucation (No	ne as reference	)						
Primary	-0.151***	0.044	0.861**	0.791- 0.937	-0.244**	0.071	0.784**	0.684-0.9	
Middle	-0.293***	0.062	0.747***	0.662- 0.842	-0.204*	0.098	0.817*	0.675- 0.988	

Table 5. The results of NPOM.

Secondary	-0.454***	0.064	0.636***	0.562- 0.72	-0.471***	0.106	0.625***	0.508- 0.769				
Higher	-0.574***	0.074	0.564***	0.488- 0.651	-0.632***	0.125	0.532***	0.417- 0.678				
Food diversi	Food diversity score (None as reference)											
Any One	-0.259***	0.06	0.773***	0.687- 0.868	-0.558***	0.091	0.573***	0.48- 0.684				
Any Two	-0.338***	0.068	0.714***	0.625- 0.815	-0.61***	0.105	0.544***	0.443- 0.668				
Any Three	-0.47***	0.083	0.626***	0.533- 0.735	-0.808***	0.135	0.447***	0.343- 0.581				
At least Four	-0.62***	0.101	0.539***	0.443- 0.655	-1.064***	0.175	0.346***	0.246- 0.487				
Size of child	at birth (Al	oove average as	reference)									
Average	0.313***	0.07	1.367***	1.193- 1.567	0.387**	0.125	1.472**	1.153- 1.88				
Below average	0.789***	0.076	2.2***	1.899- 2.55	0.977***	0.131	2.655***	2.056- 3.429				
Wealth quint	ile (Poores	t as reference)										
Poor	-0.23***	0.045	0.795***	0.729- 0.868	-0.228**	0.069	0.797**	0.697- 0.911				
Middle	-0.508***	0.053	0.602***	0.543- 0.668	-0.509***	0.084	0.602***	0.511- 0.709				
Rich	-0.53***	0.061	0.589***	0.523- 0.664	-0.586***	0.099	0.557***	0.459- 0.676				
Richest	-0.662***	0.076	0.517***	0.445- 0.599	-0.594***	0.124	0.553***	0.434- 0.705				
Divisions (Ba	ahawalpur a	as reference)										
DG Khan	-0.236***	0.066	0.79***	0.695- 0.899	-0.195	0.101	0.824	0.677- 1.003				
Faisalabad	-0.181**	0.067	0.836**	0.733- 0.952	-0.137	0.104	0.873	0.713- 1.069				
Gujranwala	-0.282***	0.068	0.755***	0.661- 0.861	-0.177	0.107	0.839	0.681- 1.034				
Lahore	-0.143*	0.07	0.868*	0.758- 0.993	-0.047	0.109	0.955	0.772- 1.182				
Multan	-0.138*	0.066	0.872*	0.766- 0.992	0.1	0.099	1.106	0.912- 1.34				
Rawalpindi	-0.562***	0.08	0.571***	0.488- 0.668	-0.202	0.123	0.818	0.644- 1.04				
Sahiwal	-0.4***	0.08	0.671***	0.574- 0.784	-0.425**	0.131	0.654**	0.506- 0.846				
Sargodha	-0.228**	0.07	0.797**	0.696- 0.913	-0.18	0.108	0.836	0.677- 1.032				

Goodness of fit (likelihood ratio); χ<sup>2</sup> (56)=1588.84; P-Value=<0.0001; Pseudo R-squared=0.0480; n=25211

 $^{\star\star\star}$  and  $^{\star\star\star}$  indicates p<0.05, p<0.01 and p<0.001, respectively.

Motivated by the overall outcomes, next we consider more sophisticated modeling scheme to handle the situation where, parallel lines assumption is violated for some variables but holds for other. The Table 6 summarizes the results for PPOM.

Comparisons									
	Not malnourished vs. (Severely and moderately (Not and moderately malnourished) vs. severely								
Covariates	Estimates	mainou Standard	Odds	95% C I	Estimates	mainou Standard	risnea Odds	95% C I	
	Lotinates	errors	ratios	(OR)	Lotimatoo	errors	ratios	(OR)	
Intercept	-0.17	0.13			-1.369***	0.221			
Child's mont	hly age (Ab	ove 23 as refere	ence)						
0-5	-0.25*	0.11	0.777*	0.63- 0.958	-0.253*	0.107	0.777*	0.63- 0.958	
06-11	-0.36***	0.09	0.7***	0.584- 0.838	-0.358***	0.093	0.7***	0.584- 0.838	
Child ill with	fever in la	st two weeks (N	o as referen	ce)				-	
Yes	0.08*	0.036	1.082*	1.008- 1.161	0.079*	0.036	1.082*	1.008- 1.161	
Child had di	arrhea in Ia	st two weeks (N	lo as referen	ice)	1		ſ	r	
Yes	0.13**	0.044	1.133**	1.041- 1.233	0.125**	0.044	1.133**	1.041- 1.233	
Received an	tenatal car	e (No as referen	ice)		1		1		
Yes	-0.24***	0.047	0.788***	0.72- 0.862	-0.24***	0.047	0.788***	0.72- 0.862	
Household h	as own agr	iculture land (N	o as referen	ce)					
Yes	-0.24***	0.035	0.79***	0.738- 0.845	-0.237***	0.035	0.79***	0.738- 0.845	
Mother's ed	ucation (No	ne as reference	)	·				·	
Primary	-0.16***	0.043	0.854***	0.785- 0.928	-0.159***	0.043	0.854***	0.785- 0.928	
Middle	-0.29***	0.061	0.751***	0.667- 0.847	-0.287***	0.061	0.751***	0.667- 0.847	
Secondary	-0.46***	0.063	0.635***	0.561- 0.718	-0.455***	0.063	0.635***	0.561- 0.718	
Higher	-0.58***	0.074	0.562***	0.487- 0.648	-0.578***	0.074	0.562***	0.487- 0.648	
Food diversi	ty score (No	one as reference	e)						
Any One	-0.26***	0.06	0.772***	0.687- 0.868	-0.556***	0.091	0.575***	0.482- 0.685	
Any Two	-0.338***	0.068	0.714***	0.625- 0.815	-0.611***	0.105	0.544***	0.443- 0.667	
Any Three		0.083	0.625***	0.532-		0.135	0.454***	0.349-	

Table 6. The results of PPOM.

	-0.471***			0.734	-0.792***			0.59
At least Four	-0.62***	0.101	0.539***	0.443- 0.656	-1.074***	0.176	0.342***	0.243- 0.482
Size of child	at birth (Al	oove average as	reference)					
Average	0.312***	0.07	1.366***	1.193- 1.566	0.404**	0.125	1.497**	1.173- 1.91
Below average	0.788***	0.076	2.199***	1.898- 2.549	0.988***	0.13	2.686***	2.082- 3.464
Wealth quint	tile (Poores	t as reference)			-			
Poor	-0.23***	0.044	0.796***	0.73- 0.867	-0.23***	0.044	0.796***	0.73- 0.867
Middle	-0.508***	0.053	0.603***	0.544- 0.668	-0.508***	0.053	0.603***	0.544- 0.668
Rich	-0.534*8*	0.061	0.587***	0.522- 0.661	-0.534***	0.061	0.587***	0.522- 0.661
Richest	-0.658***	0.076	0.519***	0.447- 0.601	-0.658***	0.076	0.519***	0.447- 0.601
Divisions (Ba	ahawalpur a	as reference)						
DG Khan	-0.229***	0.065	0.796***	0.701- 0.903	-0.229***	0.065	0.796***	0.701- 0.903
Faisalabad	-0.173**	0.066	0.842**	0.741- 0.958	-0.173**	0.066	0.842**	0.741- 0.958
Gujranwala	-0.269***	0.067	0.765***	0.671- 0.871	-0.269***	0.067	0.765***	0.671- 0.871
Lahore	-0.131	0.069	0.878	0.769- 1.004	-0.131	0.069	0.878	0.769- 1.004
Multan	-0.107	0.065	0.899	0.792- 1.021	-0.107	0.065	0.899	0.792- 1.021
Rawalpindi	-0.532***	0.08	0.588***	0.504- 0.687	-0.532***	0.08	0.588***	0.504- 0.687
Sahiwal	-0.396***	0.079	0.674***	0.577- 0.786	-0.396***	0.079	0.674***	0.577- 0.786
Sargodha	-0.22**	0.069	0.803**	0.703- 0.918	-0.22**	0.069	0.803**	0.703- 0.918
Goodness of	fit (likeliho	bod ratio); $\chi^2$ (30	6)=1553.05	; P-Value=<0.	0001; Pseu	ido R-squared=0	).0469; n=2	5211
*** and *** ir	ndicates p<	0.05, p<0.01 aı	nd p<0.001,	respectively.				

### Determinants of child nutrition status

We observed that the employed models, POM, NPOM and PPOM, reveal that all the covariates are significant predictors of malnutrition. While goodness of fit indicators is in favor of PPOM, therefore, we interpret the coefficients given in PPOM. The Table 7 shows the values of goodness of fit criteria for each model.

Table 7. Goodness of fit summary.									
Models									
Measures	POM	NPOM	PPOM						
Deviance	31623.11	31490.5	31526.28						
McFadden-R2 0.044 0.048 0.048									

Table 7.	Goodness	of fit	summary.
----------	----------	--------	----------

AIC	31683.11	31606.5	31602.28
BIC	31927.16	32078.3	31911.41

The PPOM results from Table 6 shows that the odds of falling in poor nourishment status are lowest for the children who consumed diversified diet as compared to the children who consumed none of the food item. These odds of malnutrition are 46% lower for the children who consumed at least four food groups as compared to the children whose dietary intake does not follow IYCF guidelines. This is observed when no malnutrition category is compared with moderate & severe malnutrition categories. Similarly, a consistent decrease in odds of malnutrition is observed with respect to other categories of food diversity. Such as 38% lower, 29% lower and 23% lower for the children who consumed at least any three food groups, any two food groups and anyone of the food group respectively as compared to the children with no consumption of any food. Furthermore, the odds are different when no malnutrition and moderate malnutrition categories are compared with severe malnutrition category. The results indicate that the children who consumed at least four food groups, any three food groups, any two food groups, any two food groups and anyone of the results indicate that the children who consumed at least four food groups, any three food groups, any two food groups and anyone of the food group have 66%, 55%, 47% and 43% lower risk of malnutrition, respectively, as compared to children with no consumption of any food from these complementary food groups.

Table 6 shows that the odds of falling in poor nutritional status, that is, moderate or severe malnutrition states is lowest for the children with educated mother. As observed from the Table 6, the odds of staying in extreme malnutrition state are 44%, 37%, 25% and 15% lower for children having higher, secondary, middle, and primary educated mothers, respectively, as compared to the children having uneducated mothers. Further, when the no malnutrition category is compared with moderate & severe malnutrition categories, the children with below average and average size have 1.306- and 2.199-times higher risk of falling in poor nourishment status respectively as compared to the children having above average size. On the other hand, when the no malnutrition and moderate malnutrition states are compared with severe malnutrition state, these odds remain 1.497 and 2.686 times higher. Child's health history and mother's health awareness also significantly derive the likely status of nutrition. The children who had diarrhea and fever during the last two weeks have higher odds of suffering malnourishment as compared to the children with no such type of disease. Similarly, children whose mothers received antenatal care during pregnancy have 22% lower odds of being malnourished as compared to the children from mothers who have not received the care.

While measuring the impact of wealth index quintile, odds of having worse malnutrition status are 53%, 43%, 40% and 22% lower for the children from richest, rich, middle, and poor households, respectively, as compared to poorest households. The odds of falling in worse state of malnutrition are lowest in the richest households. The households with no agriculture land have less percentage of malnourished children as compared to the households with no agriculture land. As results show that the children from these households have 23% lower odds of staying in worse state of malnutrition agriculture land. We also found significant variations in nutrition condition of children with respect to districts of the province.

### CONCLUSION

Although, Punjab is the most developed province of Pakistan still the percentage of malnourished children is very high. Despite many initiatives, there is dearth of subjective and mixed methods studies on the causes of child malnourishment in Punjab. This research advances the existing literature, mainly on two fronts. Firstly, we advocate the use of ordinal scale to quantify the nutrition stat of children in order to facilitate the identification of priority groups existent in the population. Secondly, we use IYCF proposed food diversity index, an indicator of balanced food, to explore the child malnourishment in Punjab. We also have considered additional factors related to child health that were being ignored in previous studies. Keeping in mind the nature of response variable, several contemporary modeling schemes were employed to explore the underlying ordinal structure of the nourishment status of children. Based on several good fit criteria, partial proportional odd model was selected as more sophisticated technique among many. The results of PPOM show that balance and diverse food intake has significant relationship with child's nourishment. Similarly, factors such as, mother's education and awareness, child's health history, household's socio-income status, are significantly affecting child's nourishment status.

Based on the findings of this study, we recommend three folded schemes to support the goal of achieving better child nutrition status in Punjab. We highly advocate the propagation of the use of balance diet for children. This can be persuaded through media campaign and social awareness programs. Simultaneously, mother's health awareness is an important factor. The lady health worker program can play a vital role in this regard. Thirdly, synchronized efforts are required to improve economic conditions at household level while providing enhanced opportunities of women empowerment. A more rigorous implication of the programs such as, EHSAAS and BISP while linking financial assistance to balanced diets among children is desirable.

### REFERENCES

- 1. Ananth CV, et al. Regression models for ordinal responses: A review of methods and applications. Int J Epidemiol. 1997;26:1323-1333.
- Batool S, et al. To assess the nutritional status of primary school children in an Urban school of Faisalabad. Pak J Med Health Sci A. 20122;4:160.
- 3. Bender R, et al. Using binary logistic regression models for ordinal data with non-proportional odds. J Clin Epidemiol. 1998;51:809-816.
- 4. Black RE, et al. Maternal and child under nutrition: Global and regional exposures and health consequences. Lancet. 2008;371:243-260.
- 5. Brant R, et al. Assessing proportionality in the proportional odds model for ordinal logistic regression. Biometrics. 1990;461171-1178.
- Carlson P, et al. The European health divide: A matter of financial or social capital? Soc Sci Med. 2004;59:1985– 1992.
- 7. FAO, IFAD, WFP, WHO, and UNICEF. The state of food security and nutrition in the world 2019: Safeguarding against economic slowdowns and downturns. FAO. Rome, Italy. 2019.
- 8. Fullerton AS, et al. The proportional odds with partial proportionality constraints model for ordinal response variables. Soc Sci Res. 2012;41:182–198.
- 9. Gul R, et al. Prevalence and predeterminants of malnutrition in children under 3 years of age in the two rural communities of Peshawar. Khyber Med Univ J. 2013;5:190–194.
- 10. Hoddinott J, et al. Effect of nutrition intervention during early childhood on economic productivity in Guatemalan adults. Lancet. 2008;371:411-416.
- 11. Mushtaq MU, et al. Height, weight and BMI percentiles and nutritional status relative to the international growth references among Pakistani school aged children. BMC Pediatr. 2012;12:31.
- 12. Paul GK, et al. Application of proportional odds model in identifying contributing factor of under-five child malnutrition in Bangladesh: A case study in Tangail district. J Health Res Rev. 2018;5:128.
- 13. Peterson B, et al. Partial proportional odds models for ordinal response variables. J R Stat Soc Ser C Appl Stat. 1990;39:205-217.
- Swindale A, et al. Household Dietary Diversity Score (HDDS) for measurement of household food access: Indicator guide. Food and Nutrition Technical Assistance Project, Academy for Educational Development. Washington, DC. USA. 2006.
- 15. Talukder A, et al. Factors associated with malnutrition among under five children: Illustration using Bangladesh demographic and health survey, 2014 data. Children. 2017;4:88.
- 16. Victora CG, et al. Maternal and child under nutrition: Consequences for adult health and human capital. Lancet. 2008;371:340-357.